# Multidimensional Monotonicity Discovery with mBART 

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Plan

- I) Review BART
- II) Introduce Monotone BART: mBART
- III) Monotonicity Discovery with mBART


## Part I. BART (Bayesian Additive Regression Trees)

Data: $n$ observations of $y$ and $x=\left(x_{1}, \ldots, x_{p}\right)$
Suppose: $Y=f(x)+\epsilon, \quad \epsilon$ symmetric with mean 0
Bayesian Ensemble Idea: Approximate unknown $f(x)$ by the form

$$
\begin{gathered}
f(x)=g\left(x ; \theta_{1}\right)+g\left(x ; \theta_{2}\right)+\ldots+g\left(x ; \theta_{m}\right) \\
\theta_{1}, \theta_{2}, \ldots, \theta_{m} \quad \text { iid } \sim \pi(\theta)
\end{gathered}
$$

and use the posterior of $f$ given $y$ for inference.

BART: each $g\left(x ; \theta_{j}\right)$ is a regression tree.
Key data calibration: Using $y$, set $\pi(\theta)$ so that $\operatorname{Var}(f) \approx \operatorname{Var}(y)$.

## Beginning with a Single Tree Model

Let $T$ denote the tree structure including the decision rules

Let $M=\left\{\mu_{1}, \mu_{2}, \ldots \mu_{\mathrm{b}}\right\}$ denote the set of bottom node $\mu$ 's.


Let $g(x ; T, M)$ be a regression tree function that assigns a $\mu$ value to $x$


A single tree model:

$$
Y=g(x ; T, M)+\sigma z, \quad z \sim N(0,1)
$$

## Bayesian CART: Just add a prior $\pi(M, T)$

Bayesian CART Model Search
(Chipman, George, McCulloch 1998)

$$
\pi(M, T)=\pi(M \mid T) \pi(T)
$$

$\pi(T)$ : Stochastic process to generate tree skeleton plus uniform prior on splitting variables and splitting rules.
$\pi(M \mid T):\left(\mu_{1}, \mu_{2}, \ldots, \mu_{b}\right)^{\prime} \sim N_{b}\left(0, \tau^{2} l\right)$
Closed form for $\pi(T \mid y)$ facilitates MCMC stochastic search for promising trees.

## Moving on to BART

Bayesian Additive Regression Trees
(Chipman, George, McCulloch 2010)
The BART ensemble model
$Y=g\left(x ; T_{1}, M_{1}\right)+g\left(x ; T_{2}, M_{2}\right)+\ldots+g\left(x ; T_{m}, M_{m}\right)+\sigma z, \quad z \sim N(0,1)$

Each $\left(T_{i}, M_{i}\right)$ identifies a single tree.
For each $x, Y$ is the sum of $m$ bottom node $\mu$ 's, plus noise.
Number of trees $m$ can be much larger than sample size $n$.
$g\left(x ; T_{1}, M_{1}\right), g\left(x ; T_{2}, M_{2}\right), \ldots, g\left(x ; T_{m}, M_{m}\right)$ is a highly redundant "over-complete basis" with many many parameters.

## Complete the Model with a "Regularization" Prior

$$
\pi\left(\left(T_{1}, M_{1}\right),\left(T_{2}, M_{2}\right), \ldots,\left(T_{m}, M_{m}\right), \sigma\right)
$$

$\pi$ applies the Bayesian CART prior to each $\left(T_{j}, M_{j}\right)$ independently so that:

- Each $T$ small.
- Each $\mu$ small.
- $\sigma$ will be compatible with the observed variation of $y$.

The observed variation of $y$ is used to guide hyperparameter settings for the $\mu$ and $\sigma$ priors.
$\pi$ keeps the contribution of each $g\left(x ; T_{i}, M_{i}\right)$ small, to explain only a small portion of the fit.

Build up the fit, by adding up tiny bits of fit ..


Simple prior for a complex model !!!!!!
$Y=g\left(x ; T_{1}, M_{1}\right)+g\left(x ; T_{2}, M_{2}\right)+\ldots+g\left(x ; T_{m}, M_{m}\right)+\sigma z, \quad z \sim N(0,1)$

For each $x, f(x)$ is the sum of $m$ bottom node $\mu$ 's.

$$
\begin{gathered}
\mu \sim N\left(0, \tau^{2}\right), i i d . \\
\Rightarrow f(x) \sim N\left(0, m \tau^{2}\right), \quad \forall x .
\end{gathered}
$$

When people try the $R$ package it works just by doing res = BART: :wbart (X,y).

## Connections to Other Modeling Ideas

$$
\begin{gathered}
Y=g\left(x ; T_{1}, M_{1}\right)+\ldots+g\left(x ; T_{m}, M_{m}\right)+\sigma z \\
\text { plus } \\
\pi\left(\left(T_{1}, M_{1}\right), \ldots\left(T_{m}, M_{m}\right), \sigma\right)
\end{gathered}
$$

Bayesian Nonparametrics:

- Lots of parameters (to make model flexible)
- A strong prior to shrink towards simple structure (regularization)
- BART shrinks towards additive models with some interaction


## Boosting:

- Fit becomes the cumulative effort of many weak learners


## Dynamic Random Basis Elements:

$\checkmark g\left(x ; T_{1}, M_{1}\right), \ldots, g\left(x ; T_{m}, M_{m}\right)$ are dimensionally adaptive

## A Sketch of the BART MCMC Algorithm

$$
\begin{gathered}
Y=g\left(x ; T_{1}, M_{1}\right)+\ldots+g\left(x ; T_{m}, M_{m}\right)+\sigma z \\
\text { plus } \\
\pi\left(\left(T_{1}, M_{1}\right), \ldots\left(T_{m}, M_{m}\right), \sigma\right)
\end{gathered}
$$

Bayesian Backfitting: Outer loop is a "simple" Gibbs sampler

$$
\begin{array}{r|l}
\left(T_{i}, M_{i}\right) & Y, \text { all other }\left(T_{j}, M_{j}\right), \text { and } \sigma \\
\sigma & Y,\left(T_{1}, M_{1}, \ldots, \ldots, T_{m}, M_{m}\right)
\end{array}
$$

To draw $\left(T_{i}, M_{i}\right)$ above, subtract the contributions of the other trees from both sides to get a simple one-tree model.

We integrate out $M$ to draw $T$ and then draw $M \mid T$.
... as the MCMC runs, trees in the sum will grow and shrink, swapping fit amongst them ....

## Using the MCMC Output to Draw Inference

Each iteration $d$ results in a draw from the posterior of $f$

$$
\hat{f}_{d}(\cdot)=g\left(\cdot ; T_{1 d}, M_{1 d}\right)+\cdots+g\left(\cdot ; T_{m d}, M_{m d}\right)
$$

To estimate $f(x)$ we simply average the $\hat{f}_{d}(\cdot)$ draws at $x$

Posterior uncertainty is captured by variation of the $\hat{f}_{d}(x)$ eg, $95 \%$ credible region estimated by middle $95 \%$ of values

## Out of Sample Prediction

Predictive comparisons on 42 data sets.
Data from Kim, Loh, Shih and Chaudhuri (2006) (thanks Wei-Yin Loh!)

- $p=3$ to $65, n=100$ to 7,000 .
- for each data set 20 random splits into $5 / 6$ train and $1 / 6$ test
- use 5 -fold cross-validation on train to pick hyperparameters (except BART-default!)
- gives $20 * 42=840$ out-of-sample predictions, for each prediction, divide rmse of different methods by the smallest
+ each boxplots represents 840 predictions for a method
+1.2 means you are $20 \%$ worse than the best
+ BART-cv best
+ BART-default (use default prior) does amazingly well!!



## Automatic Uncertainty Quantification

A simple simulated 1-dimensional example

95\% pointwise posterior intervals, BART


95\% pointwise posterior intervals, mBART


Note: mBART on the right plot to be introduced next

## Part II. Monotone BART - mBART

mBART: Multidimensional Monotone BART
(Chipman, George, McCulloch, Shively 2021)

Key Idea:
Approximate multivariate monotone functions by the sum of many single monotonic tree models.

This works because

1. We can easily define a notion of "monotonic" for a single tree.
2. Because trees are simple, we can construct an MCMC which respects the constraints.

So,
we can still use the BART approach
but now complex montonic functions are built as the sum of many single tree models, each of which is monotonic.

## An Example of a Monotonic Tree





Three different views of a bivariate monotonic tree.

In what sense is this tree monotonic?


A function $g$ is said to be monotonic in $x_{i}$ if for any $\delta>0$,

$$
\begin{aligned}
& g\left(x_{1}, x_{2}, \ldots, x_{i}+\delta, x_{i+1}, \ldots, x_{k} ; T, M\right) \\
& \quad \geq g\left(x_{1}, x_{2}, \ldots, x_{i}, x_{i+1}, \ldots, x_{k} ; T, M\right) .
\end{aligned}
$$

For simplicity and wlog, let's restrict attention to monotone nondecreasing functions.

Constraining a tree to be monotone is easy: we simply constrain the mean level of a node to be greater than those of its "below-neighbors", and less than those of its "above-neighbors".


The mean level of node 13 must be greater than those of 10 and 12 and less than that of node 7 .

## The mBART Prior

Recall the BART parameter

$$
\theta=\left(\left(T_{1}, M_{1}\right),\left(T_{2}, M_{2}\right), \ldots,\left(T_{m}, M_{m}\right), \sigma\right)
$$

Let $S=\left\{\theta\right.$ : every tree is monotonic in a desired subset of $\left.x_{i}^{\prime} s\right\}$
To impose the monotonicity we simply truncate the BART prior $\pi(\theta)$ to the set $S$

$$
\pi^{*}(\theta) \propto \pi(\theta) I_{S}(\theta)
$$

where $I_{S}(\theta)$ is 1 if every tree in $\theta$ is monotonic.

## A New BART MCMC "Christmas Tree" Algorithm

$$
\left.\pi\left(\left(T_{1}, M_{1}\right),\left(T_{2}, M_{2}\right), \ldots,\left(T_{m}, M_{m}\right), \sigma \mid y\right)\right)
$$

Bayesian Backfitting again: Iteratively sample each ( $T_{j}, M_{j}$ ) given $(y, \sigma)$ and other $\left(T_{j}, M_{j}\right)$ 's

Each $\left(T^{0}, M^{0}\right) \rightarrow\left(T^{1}, M^{1}\right)$ update is sampled as follows:

- Denote move as

$$
\left(T^{0}, M_{\text {Common }}^{0}, M_{\text {Old }}^{0}\right) \rightarrow\left(T^{1}, M_{\text {Common }}^{0}, M_{\text {New }}^{1}\right)
$$

- Propose $T^{*}$ via birth, death, etc.
- If M-H with $\pi(T, M \mid y)$ accepts $\left(T^{*}, M_{\text {Common }}^{0}\right)$
- Set $\left(T^{1}, M_{\text {Common }}^{1}\right)=\left(T^{*}, M_{\text {Common }}^{0}\right)$
- Sample $M_{\text {New }}^{1}$ from $\pi\left(M_{\text {New }} \mid T^{1}, M_{\text {Common }}^{1}, y\right)$

Only $M_{\text {Old }}^{0} \rightarrow M_{\text {New }}^{1}$ needs to be updated.
Works for both BART and mBART.

## the joy of working with Ed !!!!!!!



## There is the math junk and the code junk.

## But what is it like working with Ed???

```
//---------------------
// how does node n relate to this
//'a': neighbor above, 'b': neighbor below, 'd': disjoint, 'x': nothing
char tree::nhb(tree_p n, xinfo& xi, std::vector<int>& vc)
//note:
// each node is a region of the form \cap [Lv,Uv] v=0,1,\ldots,(p-1)
{
    if(this==n) return 's'; //s means self
    size_t p = xi.size(); //need to loop over the p variables
    int mL,mU,oL,oU; //my range, other range
    short nind=0;
    //first loop over all variables to see if n is disjoint from this
    for(size_t v=0;v<p;v++) {
        mL=0; oL=0;
        mU = oU = xi[v].size()-1;
        rg(v,&mL,&mU);
        n->rg(v,&oL,&oU);
        //a neighbor will be 2 away since [L,U] are the usable ones,
            //for example: 'b': [oL,oU] used cut point=C, [mL,mU]
            //would have oU,C,mL in sequence
        if(oU < mL-2 || oL > mU+2 ) return 'd';
    }
    // now loop over all variables in vc to see if n is above or below this
    for(size_t i=0;i<vc.size();i++) {
        v = vc[i];
        mL=0; oL=0;
        mU = oU = xi[v].size()-1;
        rg(v,&mL,&mU);
        n->rg(v,&oL,&oU);
        //a neighbor will be 2 away since [L,U] are the usable ones,
```

It's fun !!!!!!!!!!!!!!!!!!!!!!!!!!!!!

## Example: Product of two x's

Let's consider a very simple simulated monotone example:

$$
Y=x_{1} x_{2}+\epsilon, \quad x_{i} \sim \operatorname{Uniform}(0,1)
$$

Here is the plot of the true function $f\left(x_{1}, x_{2}\right)=x_{1} x_{2}$



First we try a single (just one tree), unconstrained tree model. Here is the graph of the fit.


The fit is not terrible, but there are some aspects of the fit which violate monotonicity.

Here is the graph of the fit with the monotone constraint:


We see that our fit is monotonic, and more representative of the true $f$.

Here is the unconstrained BART fit:


Much better (of course) but not monotone!

And, finally, the constrained BART fit:


## Not Bad!

Same method works with any number of x's!

## Example: MSE Reduction by Monotone Regularization

$$
\begin{gathered}
Y=x_{1} x_{2}^{2}+x_{3} x_{4}^{3}+x_{5}+\epsilon, \\
\epsilon \sim N\left(0, \sigma^{2}\right), x_{i} \sim \operatorname{Uniform}(0,1)
\end{gathered}
$$

For various values of $\sigma$, we simulated 5,000 observations.

RMSE improvement over unconstrained BART

| $\sigma$ | Monotone BART <br> RMSE | Unconstrained BART <br> RMSE | Percentage <br> Increase |
| :---: | :---: | :---: | :---: |
| 0.5 | 0.14 | 0.16 | $14 \%$ |
| 1.0 | 0.17 | 0.28 | $65 \%$ |



## Part III. Discovering Monotonicity with mBART

Suppose we don't know if $f(x)$ is monotone up, monotone down or even monotone at all.

Of course, a simple strategy would be simply compare the fits from BART and mBART.

Good news, we can do much better than this!
As we'll now see, mBART can be deployed to simultaneously estimate all the monotone components of $f$.

With this strategy, monotonicity can be discovered rather than imposed!

## The Monotone Decomposition of a Function

To begin simply, suppose $x$ is one-dimensional and $f$ is of bounded variation.

Any such $f$ can be uniquely written (up to an additive constant) as the sum of a monotone up function and a monotone down function

$$
f(x)=f_{\text {up }}(x)+f_{\text {down }}(x)
$$

where

- when $f(x)$ is increasing, $f_{u p}(x)$ increases at the same rate and is flat otherwise,
- when $f(x)$ is decreasing, $f_{\text {down }}(x)$ decreases at the same rate and is flat otherwise.

More precisely, when $f$ is differentiable,

$$
f_{u p}^{\prime}(x)= \begin{cases}f^{\prime}(x) & \text { when } f^{\prime}(x)>0 \\ 0 & \text { when } f^{\prime}(x) \leq 0\end{cases}
$$

and

$$
f_{\text {down }}^{\prime}(x)= \begin{cases}f^{\prime}(x) & \text { when } f^{\prime}(x)<0 \\ 0 & \text { when } f^{\prime}(x) \geq 0\end{cases}
$$

Notice the orthogonal decomposition of $f^{\prime}$

$$
f^{\prime}(x)=f_{u p}^{\prime}(x)+f_{\text {down }}^{\prime}(x)
$$

## The Monontone Discovery Strategy with mBART

Key Idea: To discover the monotone decomposition of $f$, we embed $f(x)$ as a two-dimensional function in $R^{2}$,

$$
f(x)=f^{*}(x, x)=f_{\text {up }}(x)+f_{\text {down }}(x)
$$

Letting $x_{1}=x_{2}=x$ be duplicate copies of $x$, we apply mBART to estimate $f^{*}\left(x_{1}, x_{2}\right)$

- constrained to be monotone up in the $x_{1}$ direction, and
- constrained to be monotone down in the $x_{2}$ direction.

We are effectively estimating monotone projections of $f^{*}\left(x_{1}, x_{2}\right)$ onto the $x_{1}$ and $x_{2}$ axes

- $P_{\left[x_{1}\right]} f^{*}\left(x_{1}, x_{2}\right)=f_{u p}\left(x_{1}\right)$
- $P_{\left[x_{2}\right]} f^{*}\left(x_{1}, x_{2}\right)=f_{\text {down }}\left(x_{2}\right)$

Example: Suppose $Y=x^{3}+\epsilon$.


Note that $\hat{f}_{\text {down }} \approx 0$ (the red in the right plot), as we would expect when $f$ is monotone up.

Example: Suppose $Y=x^{2}+\epsilon$.


- On the left, BART is good, but simple mBART is not.
- On the right, $\hat{f}_{\text {up }}$ and $\hat{f}_{\text {down }}$ are spot on.
- And mBARTD $=\hat{f}_{\text {up }}+\hat{f}_{\text {down }}$ seems better than BART!

Example: Suppose $Y=\sin (x)+\epsilon$.

BART and mBART

mBARTD, fup, fdown


- BART is great, but simple mBART reveals nothing.
- $\hat{f}_{u p}$ and $\hat{f}_{\text {down }}$ have discovered the monotone decomposition.
- And mBARTD $=\hat{f}_{\text {up }}+\hat{f}_{\text {down }}$ is great too.

To extend this approach to multidimensional x, we simply duplicate each and every component of $x$ !!!

## Example: House Price Data

Let's look at a simple real example where $y=$ house price, and $x=$ three characteristics of each house.

```
> head(x)
    nbhd size brick
[1,] 2 1.79 0
[2,] 2 2.03 0
[3,] 2 1.74 0
[4,] 2 1.98 0
[5,] 2 2.13 0
[6,] 1 1.78 0
> dim(x)
[1] }128 
> summary(x)
        nbhd
    Min. :1.000 Min. :1.450 Min. :0.0000
    1st Qu.:1.000 1st Qu.:1.880 1st Qu.:0.0000
    Median :2.000 Median :2.000 Median :0.0000
    Mean :1.961 Mean :2.001 Mean :0.3281
    3rd Qu.:3.000 3rd Qu.:2.140 3rd Qu.:1.0000
    Max. :3.000 Max. :2.590 Max. :1.0000
> summary(y)
        Min. 1st Qu. Median Mean 3rd Qu. Max.
        69.1 111.3 126.0 130.4 148.2 
```

$y$ : dollars (thousands).
x : nbhd (categorial), size (sq ft thousands), brick (indicator).

Call:

```
lm(formula = price ~ nbhd + size + brick, data = hdat)
```

| Residuals: |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Min | $1 Q$ | Median | $3 Q$ | Max |
| -30.049 | -8.519 | 0.137 | 7.640 | 36.912 |

Coefficients:
Estimate Std. Error t value $\operatorname{Pr}(>|t|)$

| (Intercept) | 18.725 | 10.766 | 1.739 | $0.0845 \quad$. |
| :--- | ---: | ---: | ---: | ---: | :--- |
| nbhd2 | 5.556 | 2.779 | 1.999 | $0.0478 \quad *$ |
| nbhd3 | 36.770 | 2.958 | 12.430 | $<2 e-16 \quad * * *$ |
| size | 46.109 | 5.527 | 8.342 | $1.25 \mathrm{e}-13 * * *$ |
| brickYes | 19.152 | 2.438 | 7.855 | $1.69 \mathrm{e}-12 * * *$ |

Residual standard error: 12.5 on 123 degrees of freedom
Multiple R-squared: 0.7903,Adjusted R-squared: 0.7834 F-statistic: 115.9 on 4 and 123 DF, p-value: < $2.2 \mathrm{e}-16$

If the linear model is correct, we are monotone up in all three variables.

Remark: For the linear model we have to dummy up nbhd, but for BART and mBART we can simply leave it as an ordered numerical categorical variable.

Just using $x=$ size, $y=$ price appears to be marginally increasing in size. $\left(\hat{f}_{\text {down }} \approx 0\right)$.

BART, mBART, and mBARTD

mBARTD: $y$ on ( $x$ up, $x$ down)

mBART and mBARTD seem much better than BART.

Let's now look at the effect of size conditionally on the six possible values of (nbdh, brick)




The conditionally monotone effect of size is becoming clearer!

And finally, the effect of size conditionally on the six possible values of (nbdh, brick) via $\hat{f}_{\text {up }}$ and $\hat{f}_{\text {down }}$


Price is clearly conditionally monotone up in all three variables!
By simultaneously estimating $\hat{f}_{\text {up }}+\hat{f}_{\text {down }}$, we have discovered monotonicity without any imposed assumptions!!!

## Concluding Remarks

- mBARTD $=\hat{f}_{u p}+\hat{f}_{\text {down }}$ provides an assumption free approach for the discovery of the monotone components of $f$ in multidimensional settings.
- Discovering such regions of monotonicity may of scientific interest in real applications.
- We have used informal variable selection to identify the monotone components here. More formal variable selection can be used in higher dimensional settings.
- As a doubly adaptive shape-constrained regularization approach,
- mBARTD will adapt to mBART when monotonicity is present,
- mBARTD will adapt to BART when monotonicity absent,
- mBARTD seems at least as good and maybe better, than the best of mBART and BART in general.


## Concluding Remarks

- Having come to be widely regarded as a successful approach for Bayesian machine learning, new extensions and generalizations of BART are flourishing.
- For a wonderful recent survey of many of these developments, see Hill, Linero and Murray (2020) in the Annual Review of Statistics and its Applications.


## Happy Birthday Ed!

 A true leader!