

# Outline

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# The Unknown $\mu$ Model

The model is

$$Y_i = \mu + \epsilon_i, \quad \epsilon_i \sim N(0, \sigma^2), \quad i = 1, 2, \dots, n.$$

We will assume we know  $\sigma$  and want inference for  $\mu$ .

# The Conjugate Prior

The normal is conjugate:

$$\mu \sim N(\bar{\mu}, \tau^2).$$

# The Likelihood

$$p(y_1, y_2, \dots, y_n \mid \mu) \propto e^{-\frac{1}{2\sigma^2} \sum (y_i - \mu)^2}$$

$$\sum (y_i - \mu)^2 = \sum (y_i - \bar{y})^2 + n(\mu - \bar{y})^2$$

$$L(\mu) \propto e^{-\frac{n}{2\sigma^2} (\mu - \bar{y})^2}$$

# The Posterior

$$L(\mu) \times p(\mu)$$

$$\propto e^{-\frac{n}{2\sigma^2}(\mu-\bar{y})^2} \times e^{-\frac{1}{2\tau^2}(\mu-\bar{\mu})^2}$$

$$= e^{-\frac{1}{2}(a(\mu-\bar{y})^2+b(\mu-\bar{\mu})^2)}$$

with

$$a = \frac{n}{\sigma^2}, \quad b = \frac{1}{\tau^2}$$

$$(a(\mu - \bar{y})^2 + b(\mu - \bar{\mu})^2)$$

$$= (a + b)\mu^2 - 2\mu(a\bar{y} + b\bar{\mu}) + C$$

where  $C$  does not have  $\mu$  in it.

$$= (a + b)\left(\mu^2 - 2\mu\frac{(a\bar{y} + b\bar{\mu})}{a + b}\right) + C$$

Let

$$\tilde{\mu} = \frac{(a\bar{y} + b\bar{\mu})}{a + b}.$$

$$(a(\mu - \bar{y})^2 + b(\mu - \bar{\mu})^2)$$

$$= (a + b)(\mu - \tilde{\mu})^2 + \tilde{C}$$

$$p(\mu | y) \propto e^{-\frac{a+b}{2}(\mu-\tilde{\mu})^2}$$

$$\mu | y \sim N(\tilde{\mu}, \frac{1}{a+b}).$$

with

$$a = \frac{n}{\sigma^2}, \quad b = \frac{1}{\tau^2}, \quad \tilde{\mu} = \frac{(a\bar{y} + b\bar{\mu})}{a+b}.$$



# Comments

## Note:

You can think of  $a$  as the weight given to the data and  $b$  as the weight given to the prior.

As  $n$  gets big the data dominates, as  $\tau$  gets small the prior dominates.

$b$  is called the prior *precision*.

The posterior mean is a weighted combination of the sample mean and the prior mean.

Note:

If you let  $\tau \rightarrow \infty$  to get

$$\mu \sim N(\bar{y}, \frac{\sigma^2}{n})$$

which looks like the frequentist approach.

If you use the “improper prior”  $p(\mu) \propto 1$  you get the above.

This prior is often describe as “ non-informative”.

Don't do this!

## Note:

We often use the posterior mean as our estimate of  $\mu$ .

This is optimal under squared error loss.

For finite  $\tau$  we “shrink” the posterior from the sample mean towards the prior mean, where the amount of shrinkage is determined by  $\tau$ .

$$L(\mu) \times p(\mu)$$

$$\propto e^{-\frac{n}{2\sigma^2}(\mu - \bar{y})^2} \times e^{-\frac{1}{2\tau^2}(\mu - \bar{\mu})^2}$$

If we take the log we get

$$\log(p(\mu | y)) = -\frac{n}{2\sigma^2}(\mu - \bar{y})^2 - \frac{1}{2\tau^2}(\mu - \bar{\mu})^2$$

If we set  $\bar{\mu} = 0$ , the second term is called a “penalty term” by data-miners.

Note:

For finite  $\tau$ , the posterior mean is always a biased estimator.

Good.

# Prediction

Let  $X$  denote the “next  $X$ ”.

What is the predictive distribution of  $X$ ?

Have

$$X \mid \mu \sim N(\mu, \sigma^2).$$

$$\mu \sim N(\bar{\mu}, \tau^2),$$

where, again,  $\bar{\mu}$  and  $\tau$  could correspond to the prior or posterior versions.

then

$$Y = \mu + \epsilon, \epsilon \sim N(0, \sigma^2)$$

$$\sim N(\bar{\mu}, \sigma^2 + \tau^2).$$