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## Markov Chains

Let  $\{Y_t\}$   $t=0,1,2, \dots$

be a sequence of random variables.

$\{Y_t\}$  is a Markov chain if

$$(a) \quad Y_{t+1} | Y_{t-1}, Y_{t-2}, \dots, Y_0 \\ = Y_{t+1} | Y_{t-1}$$

$$(b) \quad P_t(Y_{t+1} | Y_{t-1}) = P(Y_{t+1} | Y_{t-1})$$

so "p" is the distribution of the next  $y$  given the previous  $y$ .

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Example

$$Y_t \in \{1, 2\}$$

$$P(Y_t = i | Y_{t-1} = j) = P_{ji}$$

$$P = [P_{ji}] = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}$$

" $P_{ji} = P_{j \rightarrow i}$ "

$$P(\text{next is 1} | \text{at 1}) = \frac{1}{2}$$

$$P(\text{next is 2} | \text{at 2}) = \frac{3}{4}$$

Example AR(1)

$$Y_t = \alpha + \beta Y_{t-1} + \varepsilon_t$$

$$\varepsilon_t \sim N(0, \sigma^2) \text{ i.i.d.}$$

# Stationary Distribution

Suppose  $y_0 \sim \pi$

$$p(y_1, y_0) = \pi(y_0) p(y_1 | y_0)$$

if  $y_1 \sim \pi$  we say

$\pi$  is the stationary distribution of the Markov chain given by  $p(y_t | y_{t-1})$ .

(a) draw  $y_0 \sim \pi$

(b) draw  $y_1 \sim p(y_1 | y_0)$

if  $y_1 \sim \pi$ ,  $\pi$  is stationary.

Example

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}$$

$$\pi = \left[ \frac{1}{3}, \frac{2}{3} \right]$$

$$P(Y_1=1) = P(Y_0=1) \cdot P(Y_1=1 | Y_0=1) \\ + P(Y_0=2) \cdot P(Y_1=1 | Y_0=2)$$

$$= \left(\frac{1}{3}\right)\left(\frac{1}{2}\right) + \left(\frac{2}{3}\right)\left(\frac{1}{4}\right)$$

$$= \left(\frac{1}{3}\right)\left(\frac{1}{2}\right) + \left(\frac{1}{3}\right)\left(\frac{2}{4}\right) = \frac{1}{3}$$

$$Y_1 \sim \left[ \frac{1}{3}, \frac{2}{3} \right] = \pi$$

Note

$$\pi_1 = \pi P$$

$$Y_1 \sim \pi_1$$

$$\left[ \frac{1}{3}, \frac{2}{3} \right] \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix} = \left[ \frac{1}{3}, \frac{2}{3} \right]$$

(5)

Note

$$\underline{I} \quad Y_t \sim \pi_t \quad Y_0 \sim \pi$$

$$\pi_t = \pi P^t$$

e.g.  $\pi_2 = \begin{pmatrix} \pi P \\ \pi_1 \end{pmatrix} (P) = \pi P^2$

## Example

$$Y_t = \alpha + \beta Y_{t-1} + \varepsilon_t$$

$$\varepsilon_t \sim N(0, 1)$$

$Y \sim N\left(\frac{\alpha}{1-\beta}, \frac{\alpha^2}{1-\beta^2}\right)$  is stationary.

Suppose  $Y_0 \sim N\left(\frac{\alpha}{1-\beta}, \frac{\alpha^2}{1-\beta^2}\right)$

$Y_1 = \alpha + \beta Y_0 + \varepsilon$  is normal.

$$\begin{aligned} E(Y_1) &= \alpha + \beta \left[ \frac{\alpha}{1-\beta} \right] \\ &= \frac{\alpha(1-\beta) + \beta\alpha}{1-\beta} = \frac{\alpha}{1-\beta} \end{aligned}$$

$$\begin{aligned} \text{Var}(Y_1) &= \beta^2 \frac{\alpha^2}{1-\beta^2} + \alpha^2 \\ &= \frac{\alpha^2}{1-\beta^2} \end{aligned}$$

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## Two Basic Results

$$\text{Let } Y_0 \sim P_0 \quad P(Y_t | Y_{t-1})$$

$$Y_t \sim P_t$$

A If  $\pi$  is stationary,

$$P_t \rightarrow \pi \quad \text{as } t \rightarrow \infty$$

(no matter what  $P_0$  is).

Example

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}$$

$$P^t \rightarrow \begin{bmatrix} \pi \\ \pi \end{bmatrix} = \begin{bmatrix} \frac{1}{3}, & \frac{2}{3} \\ \frac{1}{3}, & \frac{2}{3} \end{bmatrix}$$

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Suppose  $\pi$  is the stationary distribution

We want  $E(f(Y))$

$Y \sim \pi$

Then  $E(f(Y)) \approx \frac{1}{T} \sum_{t=1}^T f(Y_t)$

where  $\{Y_t\}$  are generated using the Markov chain.



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Example

$$f(y) = \chi_A(y) = \begin{cases} 1 & y \in A \\ 0 & \text{else} \end{cases}$$

$$P(Y \in A) = E[f(Y)]$$

$$\approx \int_0^1 \chi_A(x) dx$$

So

Histogram of  $\{Y_t\}$   
is like  $\pi$ !

In short

MC:

$$Y_t \sim \pi \quad \text{iid}$$

$$\{Y_t\} \sim \pi$$

MCMC:

$\{Y_t\}$  Markov.

$\pi$  is stationary

$$\{Y_t\} \sim \pi$$

Note

If  $Y_0 \sim \pi_0$

is very different from

$\pi$  then it may

"take a while" for  
the chain to "burn in"

so we may drop some  
initial draws.

$$E(f(Y)) \approx \frac{1}{T - T_0} \sum_{t=T_0+1}^T f(Y_t)$$